

C ONVERGENCE

STAT 510

WHAT HAPPENS WITH MORE DATA ?

MAIN IDEAS

- LAW OF LARGE NUMBERS

HIGH PROBABILITY THAT \bar{X}_n CLOSE TO μ

- CENTRAL LIMIT THEOREM

\bar{X}_n IS APPROX NORMAL

WHEN n "LARGE"

DEFN

GIVEN A SEQUENCE OF RVS X_1, X_2, \dots

AND ANOTHER RV X

WE SAY X_n CONVERGES IN PROBABILITY TO X ,

$$X_n \xrightarrow{P} X$$

IF FOR ANY $\epsilon > 0$

$$P(|X_n - X| > \epsilon) \rightarrow 0 \text{ AS } n \rightarrow \infty$$

DEFN

GIVEN A SEQUENCE OF RVs X_1, X_2, \dots

AND ANOTHER RV X

LET F_n BE THE CDF OF X_n AND F BE THE CDF OF X

WE SAY X_n CONVERGES IN DISTRIBUTION TO X ,

$$X_n \xrightarrow{d} X$$

IF

$$\lim_{n \rightarrow \infty} F_n(t) = F(t)$$

FOR ALL t WHERE F IS CONTINUOUS

POINT MASS DISTRIBUTION

$$P(X = c) = 1$$

$$X_n \xrightarrow{P} X$$

$$X_n \xrightarrow{P} c$$

$$X_n \xrightarrow{d} X$$

$$X_n \xrightarrow{d} c$$

EXAMPLE

$$X_n \sim N(0, 1_n) \quad X_n \xrightarrow{?} 0$$

In dist? Let F be CDF for point mass @ 0.

$$F(x) = \begin{cases} 0 & x < 0 \\ 1 & x \geq 0 \end{cases} \quad Z = \sqrt{n} X_n \sim N(0, 1)$$

$$\epsilon < 0 \quad F_n(+)=P[X_n \leq +] = P[\sqrt{n} X_n \leq \sqrt{n}+] = P[Z \leq \sqrt{n}+] \rightarrow 0$$

$$\epsilon > 0 \quad F_n(+)=P[Z \leq \sqrt{n}+] \rightarrow 1$$

$$F_n(+)\rightarrow F(+)\quad \underline{\epsilon \neq 0}$$

$$\epsilon = 0 \quad F_n(0) = 1/2 \neq F(0) = 1$$

not needed,

$$X_n \xrightarrow{?} 0$$

EXAMPLE

$$X_n \sim N(0, V_n) \quad X_n \xrightarrow{?} 0$$

IN PROB ? LET $\epsilon > 0$

$$P(|X_n| > \epsilon) = P(|X_n|^2 > \epsilon^2)$$

$$\leq \frac{\mathbb{E}[X_n^2]}{\epsilon^2} = \frac{\mathbb{V}[X_n] + \mathbb{E}[X_n]^2}{\epsilon^2} = \frac{V_n}{\epsilon^2} \rightarrow 0$$



$$X_n - 0$$

$$X_n \xrightarrow{P} 0$$

$$X_n \xrightarrow{d} X \neq X_n \xrightarrow{P} X$$

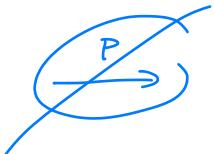
$E[X_n] = -1 \cdot 0 = 0$
 $V[X_n] = (-1)^2 \cdot 1 = 1 \quad \checkmark$

$$X \sim N(0, 1) \quad X_n = -X \quad \text{for } n = 1, 2, 3, \dots$$

← SAME DISTRIBUTION →
 $N(0, 1)$

$$\lim_{n \rightarrow \infty} F_n(x) = F(x) \quad X_n \xrightarrow{d} X$$

$$P(|X_n - x| > \epsilon) = P(|2x| > \epsilon) = P[|x| > \frac{\epsilon}{2}] \neq 0$$



DEFN

X_n CONVERGES TO X IN QUADRATIC MEAN

$$X_n \xrightarrow{qm} X$$

IF $\mathbb{E}[(X_n - X)^2] \rightarrow 0$ AS $n \rightarrow \infty$

T_{Hm}

• $X_n \xrightarrow{q_m} X \Rightarrow X_n \xrightarrow{P} X$

• $X_n \xrightarrow{P} X \Rightarrow X_n \xrightarrow{D} X$

• IF $P[X = c] = 1$ AND $X_n \xrightarrow{P} X$

THEN $X_n \xrightarrow{P} X$

REVERSES NOT TRUE IN GENERAL

THM

X_n, X, Y_n, Y Rus , \circ CONTINUOUS FUNCTION

• $X_n \xrightarrow{P} X, Y_n \xrightarrow{P} Y \Rightarrow X_n + Y_n \xrightarrow{P} X + Y$

• $X_n \xrightarrow{J} X, Y_n \xrightarrow{J} c \Rightarrow \underline{X_n + Y_n} \xrightarrow{J} \underline{X_n + c}$

• $X_n \xrightarrow{P} X, Y_n \xrightarrow{P} Y \Rightarrow X_n Y_n \xrightarrow{P} XY$

• $X_n \xrightarrow{J} X, Y_n \xrightarrow{J} c \Rightarrow \underline{cX_n} \xrightarrow{J} \underline{cX}$

• $X_n \xrightarrow{P} X \Rightarrow g(X_n) \xrightarrow{P} g(X)$

• $X_n \xrightarrow{J} X \Rightarrow g(X_n) \xrightarrow{J} g(X)$

SLUTSKY'S
THM

CONTINUOUS
MAPPING
THM

T HM

WEAK LAW OF LARGE NUMBERS (WLLN)

GIVEN X_1, X_2, \dots, X_n iid $E[X_i] = \mu$

$$\bar{X}_n \xrightarrow{P} \mu$$

PROOF

Assume $\sigma^2 < \infty$.

$$P(|\bar{X}_n - \mu| > \epsilon) \leq \frac{\mathbb{V}[\bar{X}_n]}{\epsilon^2} = \frac{\sigma^2}{n \epsilon^2} \rightarrow 0 \quad \checkmark$$

T hm

CENTRAL LIMIT THEOREM (CLT)
"CLASSICAL"

X_1, X_2, \dots, X_n iid $E[X_i] = \mu$, $V[X_i] = \sigma^2 < \infty$

DEFINE $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$

THEN

$$Z_n = \frac{(\bar{X}_n - \mu)}{\sqrt{V[\bar{X}_n]}} = \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{D} Z$$

$$Z \sim N(0, 1)$$

APPROXIMATING PROBABILITIES

STATEMENTS LIKE $P(\bar{X}_n > c) \approx \text{Normal}$

→ "FOR ANY DISTRIBUTION OF X :

→ "WHEN n LARGE"

ALTERNATIVE Forms

$$Z_n \approx N(0, 1)$$

→ $\bar{X}_n \approx N(\mu, \sigma^2/n)$

$$\bar{X}_n - \mu \approx N(0, \sigma^2/n)$$

$$n(\bar{X}_n - \mu) \approx N(0, \sigma^2)$$

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \approx N(0, 1)$$

$\bar{X}_n \xrightarrow{J} N(\mu, \sigma^2/n)$

↑
CAN'T SAY
BECAUSE OF THIS

CUT WITH S_n

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{S_n} \xrightarrow{D} N(0, 1)$$

PROOF BY EXERCISE!

T_{HM}

BERRY - ESSEN INEQUALITY

GIVEN $E[|X_1|^3] < \infty$ THEN

$$\sup_z |P(Z_n \leq z) - \Phi(z)| \leq \frac{33}{4} \frac{E[|X_1 - \mu|^3]}{\sqrt{n} r^3}$$

\nearrow
CDF OF $N(0,1)$

MAYBE TRY SIMULATIONS....

T Hm

DELTA METHOD

GIVEN $\frac{\sqrt{n} (Y_n - \mu)}{\sigma} \xrightarrow{D} N(0, 1)$

AND g IS DIFFERENTIABLE WITH $g'(\mu) \neq 0$

THEN $\frac{\sqrt{n} (g(Y_n) - g(\mu))}{|g'(\mu)| \sigma} \xrightarrow{D} N(0, 1)$

$$\text{IF } Y_n \sim N(\mu, \sigma^2/n)$$

$$\text{THEN } g(Y_n) \approx N\left(g(\mu), (g'(\mu))^2 \sigma^2/n\right)$$